

# Online Appendix for: “Beyond Income: Health, Wealth, and Racial Welfare Gaps Among Older Americans”

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## A Multiple imputation of consumption and other missing data

We followed the procedure of Miller and Bairoliya (2023) to impute consumption and other missing HRS data. This procedure uses the bootstrapping approach for cross-sectional time-series data proposed by Honaker and King (2010) and was implemented through the Amelia II software program (Honaker et al., 2011). We used the program to produce twelve complete datasets without missing data and all analyses were conducted on each dataset then combined into a single estimate. We followed Miller and Bairoliya (2023) in selecting the following variables for the imputation model: number of household members, age, aged squared, cubed root of total wealth, log household income, and dummy indicators for cohort, labor force status, gender, race, education, marital status, census division, 1980 census occupation code for longest reported tenure, self-reported health, ADLs, and eight doctor diagnosed health conditions. Additionally, our model accounted for retirement, hours worked, and an alternate measure of consumption that included health spending. To account for the time-series nature of the data, we included lags and leads of consumption, wealth, income, and hours worked in our imputation model. Interested readers can consult the appendix in Miller and Bairoliya (2023) for further details and diagnostic tests that indicate the procedure’s effectiveness in imputing missing data in the HRS dataset.

## B Forecasting model

In this section we detail our estimation and simulation procedures.

### B.1 Higher order lags

In order to avoid autocorrelation within the structural error terms of the model, it may be necessary to consider additional outcome lags. An extension of the VAR(1) model to higher orders is straightforward, as seen with the following VAR(2) version of our model:

$$AY_{it} = BY_{it-1} + DY_{it-2} + CX_{it} + \varepsilon_{it},$$

with the block matrix form of  $DY_{it-2}$  given by:

$$\begin{bmatrix} D_{11} & D_{12} & D_{13} & D_{14} & D_{15} \\ D_{21} & d_{22} & d_{23} & d_{24} & d_{25} \\ D_{31} & d_{32} & d_{33} & d_{34} & d_{35} \\ D_{41} & d_{42} & d_{43} & d_{44} & d_{45} \\ D_{51} & d_{52} & d_{53} & d_{54} & d_{55} \end{bmatrix} \begin{bmatrix} M_{it-2} \\ s_{it-2} \\ r_{it-2} \\ c_{it-2} \\ w_{it-2} \end{bmatrix}.$$

For example, the coefficient vector  $D_{51}$  in this model allows the second lag of the morbidity state vector to have a direct effect on current wealth. We can also shut down any specific lag by setting the appropriate coefficient to zero. For example, excluding the second lag of self-rated health on wealth simply implies setting  $d_{52} = 0$ .

## B.2 Estimation

The forecasting model was estimated using a pooled sample of individuals born before 1966 who were aged fifty or older at the time of the survey. The sample included 39,635 unique individuals and a total of 262,736 individual-year observations. Table 1 presents descriptive statistics for each cohort in the HRS.

Table 1: Estimation sample descriptive statistics by cohort

	AHEAD	CODA	EHRH	LHRS	WB	EBB	MBB	LBB
Individuals	7,651	4,146	5,258	5,135	3,555	4,640	4,916	4,334
Observations	36,872	29,080	47,908	50,213	31,941	30,403	24,348	11,971
Age (mean)	81.83	75.78	68.55	63.92	61.95	60.13	57.03	54.38
Hypertension (%)	54.74	57.86	54.70	52.39	51.98	52.52	50.59	49.16
Diabetes (%)	15.48	19.03	19.77	19.10	19.76	22.24	21.90	22.26
Cancer (%)	16.82	18.01	14.73	11.75	11.78	9.57	8.56	7.56
Lung disease (%)	9.43	10.24	9.81	8.89	7.82	7.57	8.29	8.23
Heart disease (%)	35.35	31.56	24.08	20.53	18.32	16.30	13.91	11.98
Stroke (%)	15.37	12.47	7.85	6.55	6.35	5.47	4.98	4.65
Psyche problem (%)	11.84	11.85	11.38	13.22	17.55	19.89	20.80	21.94
Arthritis (%)	56.03	60.76	58.31	54.30	54.07	49.60	43.33	37.57
Difficulty with ADLs (%)	40.54	29.79	25.21	23.09	23.35	23.31	21.75	18.45
Self-rated health (%)								
Poor	14.18	10.31	9.23	7.73	6.42	7.25	6.84	7.06
Fair	25.75	22.00	19.60	19.17	17.22	19.87	21.86	22.82
Good	30.90	32.21	31.84	31.30	31.27	30.99	32.05	32.05
Very good	21.41	26.40	28.02	28.83	32.04	30.50	29.30	27.94
Excellent	7.76	9.08	11.31	12.96	13.04	11.38	9.95	10.13
Retired (%)	95.51	91.77	78.47	67.29	63.28	56.51	48.68	39.79
Annual consumption (\$1000s, mean)	22.39	24.76	24.67	25.80	26.41	23.27	19.95	18.95
Male (%)	37.41	46.33	44.76	45.18	37.62	42.33	42.48	41.12
Education (%)								
<HS	41.49	31.76	30.76	28.01	20.97	20.00	21.99	21.99
HS	29.75	31.85	32.92	33.10	30.94	24.46	24.84	23.92
Some College	16.46	17.94	18.70	20.57	24.41	28.54	29.87	29.05
College	12.30	18.46	17.62	18.31	23.68	27.00	23.31	25.04
Race (%)								
White	80.91	83.44	75.67	72.97	76.02	60.58	52.65	47.68
Black	12.92	9.70	16.38	16.00	15.09	22.27	26.97	29.04
Hispanic	6.17	6.85	7.96	11.03	8.89	17.15	20.38	23.28

Notes: Children of the Depression denoted by CODA, War Babies by WB, early Baby Boomers by BB, mid Baby Boomers by MBB, and late Baby Boomers by LBB. Consumption is reported in real 2010 dollars. Source: HRS.

We estimate each block in our model separately as there is no simultaneity across blocks. As shown in the methods section of the paper, the consumption and wealth blocks only consist of one equation which follows a standard linear dynamic panel data model with lagged dependent variables and fixed effects at the individual level. We estimate these equations with OLS, but to avoid the Nickell (1981) bias that OLS can generate for this kind of model, we use the Everaert and Pozzi (2007) bootstrap-based method.<sup>1</sup> By including a single period lag of retirement and health

<sup>1</sup>We implement the bootstrap with De Vos et al. (2015) Stata routine *xtbcfe*.

on consumption, and two lags of consumption on itself, we ensure that shocks are not serially correlated in the consumption equation. Similar lags are included in the wealth equation. We also use a VAR(2) system in the retirement, health, and survival equations to maintain consistency with the consumption and wealth models. The self-rated health equation is estimated independently of other VAR blocks via maximum likelihood, while the retirement and mortality equations are estimated independently using standard probit regressions.<sup>2</sup>

Finally, we estimate the morbidity block, which we model as a multivariate probit with correlated shocks. To estimate this model, we use a chain of bivariate probit estimators suggested by Mullahy (2016) because of the large number of outcomes and observations in the HRS. While this approach allows for consistent estimation via maximum likelihood, a potential issue arises due to the absorbing nature of morbidity states. This means, for example, diagnosed hypertension at time  $t$  perfectly predicts hypertension at time  $t + 1$  and we have quasi-complete separation. In a univariate probit model, we could condition on not being diagnosed with the morbidity at time  $t$  to solve this issue, but in the bivariate probit this is not possible. Thus, we constrain the infinite coefficients to large values in the bivariate probit, but this does not affect the likelihood or estimates of remaining (non-infinite) coefficients.

The full set of estimation results are shown in Tables 3-5.

### B.3 Simulations

We used the estimated panel VAR model to construct the expected remaining lifetime utility for a subset of sixty-year-olds from the HRS. Analyses are limited to the EHRS, LHRS, War Babies, and early and mid Baby Boomer cohorts as simulations require data at age fifty-eight and sixty as "initial" conditions. The HRS provides respondent-level analysis weights for each wave, designed to create representative cohort samples of the non-institutionalized US population. We used base year weights corresponding to when the cohort was approximately age sixty to analyze the welfare distribution. Specifically, we used the 1996 analysis weights for the EHRS, 2000 for the LHRS, 2006 for War Babies, 2012 for early Baby Boomers, and 2018 for mid Baby Boomers. As any missing data was imputed among respondents, no individuals were removed from the simulation due to missing item response. However, individuals were removed if they were not surveyed at ages 58-59 and 60-61. For example, any EHRS cohort member interviewed at age 60 in 1996 but missing from the previous survey round would be excluded from the simulation sample but included in the 1996 nationally representative sample. Table 2 provides a comparison of time invariant characteristics between the weighted representative sample and the sample used in our simulations after dropping these missing cases. The simulation sample was slightly more female, educated, and white in comparison to the representative sample, but the differences were minor and generally consistent across all cohorts.

Using age sixty data as initial ( $t = 0$ ) conditions<sup>3</sup>, we simulate the remaining life outcomes for each individual ( $i$ ) as follows:

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<sup>2</sup>There is no incidental parameters or initial conditions problem in these models as there is no permanent unobserved heterogeneity or serial correlation. The standard (ordered) probit estimator is consistent and provides asymptotically valid test statistics and standard errors.

<sup>3</sup>Initial conditions also include unobserved endowments  $\hat{\pi}$  estimated using the prediction method of De Vos et al. (2015).

Table 2: Representative and simulation sample comparison

	EHRS		LHRS		WB		EBB		MBB	
	Rep	Sim	Rep	Sim	Rep	Sim	Rep	Sim	Rep	Sim
	0	1	2	3	4	5	6	7	8	9
Individuals	3,094	3,031	3,808	3,533	2,630	2,509	3,727	3,138	3,249	3,042
Male (%)	47.13	46.30	46.82	46.64	47.80	47.82	47.63	47.16	47.64	47.62
Education (%)										
<HS	29.32	29.16	25.37	25.44	18.56	18.21	13.73	12.42	14.53	14.41
HS	33.77	33.92	32.35	32.63	30.61	30.53	23.29	23.51	22.18	22.25
Some College	19.30	19.28	21.65	21.44	24.22	24.31	29.31	28.97	30.79	30.44
College	17.61	17.64	20.63	20.50	26.62	26.94	33.67	35.10	32.50	32.90
Race (%)										
White	83.37	83.89	81.55	81.92	82.46	83.06	79.55	82.99	76.38	76.47
Black	10.55	10.43	10.12	10.08	9.64	9.13	11.23	9.12	12.20	12.01
Hispanic	6.08	5.68	8.34	8.00	7.90	7.82	9.22	7.89	11.42	11.52

Notes: War Babies denoted by WB and Baby Boomers by BB. EHRS cohort includes those under age 60 in 1992. "Rep" indicates representative sample based on HRS respondent analysis weights. "Sim" indicates simulation sample weighted by the same analysis weights.

1. Survival shock  $u_{i1}$  is drawn and survival to time  $t = 1$  (age 62) is determined according to the mortality equation. If individual survives, move to step two.
2. Morbidity shock vector  $e_{i1}$  is drawn from a standard multivariate normal distribution with estimated covariance matrix  $\Sigma$  (see Table 5). This shock vector along with the model outlined in the methods section is used to compute simulated age 62 morbidity vector  $M_{i1}$ .
3. Given age 62 morbidities ( $M_{i1}$ ), general health shock  $\varepsilon_{2,i1}$  is drawn and age 62 self-rated health ( $s_{i1}$ ) is computed.
4. Given age 62 self-rated health ( $s_{i1}$ ) and morbidities ( $M_{i1}$ ), retirement shock  $\varepsilon_{3,i1}$  is drawn to determine age 62 retirement ( $r_{i1}$ ).
5. Given age 62 retirement, self-rated health, and morbidities ( $r_{i1}, s_{i1}, M_{i1}$ ), consumption shock  $\varepsilon_{4,i1}$  is drawn to determine age 62 consumption ( $c_{i1}$ ).<sup>4</sup>
6. Given all other age 62 outcomes ( $c_{i1}, r_{i1}, s_{i1}, M_{i1}$ ), wealth shock  $\varepsilon_{5,i1}$  is drawn to determine age 62 wealth ( $w_{i1}$ ).
7. Steps 1-6 are repeated for  $t = 2, 3, \dots$  until death or  $t = 30$  (age 120).
8. Steps 1-7 are repeated 5,000 times for each individual.

Figures 1-4 show a comparison between the average simulated life-cycle profiles and those constructed from available data by race for the EHRS cohort. The simulations closely match the

<sup>4</sup> $\varepsilon_4$  is drawn from the normal distribution with mean zero and standard deviation determined to match the empirical error distribution of each cohort. Specifically, standard deviations used for EHRS, LHRS, WB, and BB cohorts are 0.49, 0.48, 0.48, and 0.40. Clustering by cohort provides a slightly better fit to the data.

available aggregated data, indicating that our life-cycle dynamics model is a reasonable approximation of the underlying data generating processes. Note that the data and simulations are the same at age 60 by construction. However, the simulations match the data quite well even up to 24 years later, when the EHRS cohort reaches age 84.

To further demonstrate the accuracy of our model, we compare consumption and health utility means and standard deviations of the data with simulated life-cycle profiles for each birth cohort in Figures 5-6. The simulations match the data well across birth cohorts, further highlighting the advantages of using the VAR approach to forecast joint dynamics accurately.

## B.4 Figures and Tables

Table 3: Model estimates for ADLs, self-rated health, retirement, consumption, and mortality

Variable	ADLs		Self-rated health		Retirement		Consumption		Mortality	
	Coeff	SE	Coeff	SE	Coeff	SE	Coeff	SE	SE	SE
Hyper			-0.280	0.014	0.072	0.034	0.001	0.015	0.104	0.025
Diab			-0.257	0.017	0.060	0.043	-0.003	0.017	0.100	0.030
Cancer			-0.684	0.018	0.186	0.048	0.035	0.017	0.653	0.024
Lung			-0.468	0.021	0.187	0.065	-0.012	0.024	0.402	0.029
Heart			-0.484	0.015	0.103	0.043	-0.004	0.014	0.190	0.023
Stroke			-0.491	0.020	0.484	0.069	-0.059	0.020	0.238	0.027
Psych			-0.416	0.020	0.377	0.055	-0.046	0.022	0.225	0.028
Arthritis			-0.223	0.013	0.031	0.032	0.013	0.012	-0.030	0.023
ADL			-0.669	0.012	0.401	0.036	-0.049	0.014	0.349	0.018
Health 2					-0.563	0.044	0.051	0.016	-0.336	0.016
Health 3					-0.710	0.045	0.068	0.018	-0.536	0.017
Health 4					-0.725	0.046	0.092	0.016	-0.654	0.021
Health 5 (best)					-0.714	0.051	0.121	0.019	-0.651	0.030
Lag Hyper	0.028	0.029	0.156	0.018	-0.017	0.046	-0.002	0.013	-0.048	0.025
Lag Diab	0.102	0.036	0.103	0.023	-0.017	0.061	-0.007	0.015	0.057	0.031
Lag Cancer	0.051	0.040	0.526	0.026	-0.129	0.075	-0.017	0.018	-0.444	0.026
Lag Lung	0.188	0.045	0.206	0.031	0.023	0.097	-0.003	0.018	-0.126	0.031
Lag Heart	0.070	0.031	0.284	0.021	-0.137	0.064	0.000	0.015	-0.036	0.023
Lag Stroke	0.374	0.043	0.373	0.029	-0.254	0.116	-0.008	0.020	-0.055	0.029
Lag Psych	0.335	0.041	0.231	0.028	-0.134	0.081	0.018	0.021	-0.142	0.029
Lag Arthritis	0.227	0.024	0.114	0.017	0.046	0.042	-0.003	0.011	-0.082	0.022
Lag ADL			0.335	0.017	-0.207	0.054	0.004	0.012	-0.124	0.018
Lag Health 2	-0.246	0.029	0.634	0.013	-0.006	0.057	0.014	0.011	-0.061	0.017
Lag Health 3	-0.490	0.029	1.147	0.014	-0.044	0.058	0.017	0.014	-0.103	0.019
Lag Health 4	-0.665	0.032	1.689	0.015	-0.080	0.060	0.017	0.015	-0.142	0.022
Lag Health 5	-0.751	0.038	2.312	0.017	-0.077	0.063	0.023	0.016	-0.158	0.029
Time	-0.048	0.006	0.019	0.003	-0.000	0.009	0.002	0.009	-0.013	0.004
2008+	0.025	0.023	0.003	0.011	-0.077	0.031	-0.056	0.010	0.044	0.020
CODA	0.094	0.030	0.019	0.015	0.052	0.076			-0.012	0.023
Early HRS	0.132	0.042	0.020	0.021	0.054	0.087			-0.045	0.031
Late HRS	0.148	0.054	0.015	0.026	-0.032	0.100			-0.064	0.040
War Babies	0.187	0.067	0.002	0.032	0.005	0.116			-0.136	0.050
Early Boomers	0.290	0.081	-0.055	0.039	0.004	0.136			-0.161	0.061
Mid Boomers	0.324	0.096	-0.096	0.046	-0.089	0.154			-0.205	0.073
Late Boomers	0.390	0.115	-0.089	0.055	0.033	0.175			-0.325	0.098
Black	0.090	0.017	-0.067	0.008	0.051	0.021			0.027	0.015
Hispanic	0.048	0.022	-0.128	0.010	0.012	0.026			-0.198	0.020
Female	-0.004	0.013	0.037	0.006	0.124	0.016			-0.212	0.012
HS grad	-0.083	0.016	0.073	0.007	-0.025	0.022			0.011	0.013
Some college	-0.032	0.018	0.106	0.008	-0.048	0.024			0.003	0.016
College grad	-0.091	0.021	0.186	0.010	-0.059	0.026			-0.034	0.019
Retired							-0.036	0.013	0.193	0.030
Lag Retired	0.123	0.027	-0.023	0.012			-0.034	0.014	-0.011	0.026
Lag2 Retired	-0.016	0.025	-0.015	0.012						
Lag Con							0.164	0.004		
Lag2 Con							0.079	0.004		
Constant	-0.916	0.072			-0.898	0.178			-1.659	0.242

Notes: Dependent variable across columns. Multivariate probit results reported for ADLs as dependent outcome. Standard (ordered) probit results reported for self-rated health, mortality, and retirement as dependant outcomes. Linear dynamic panel estimates reported for consumption as outcome. All regressions also include dummies for age. Regressions for ADLs, self-rated health, mortality, and retirement also include dummies for occupation and census division. Regressions for ADLs and self-rated health also includes second lag for all health outcomes.



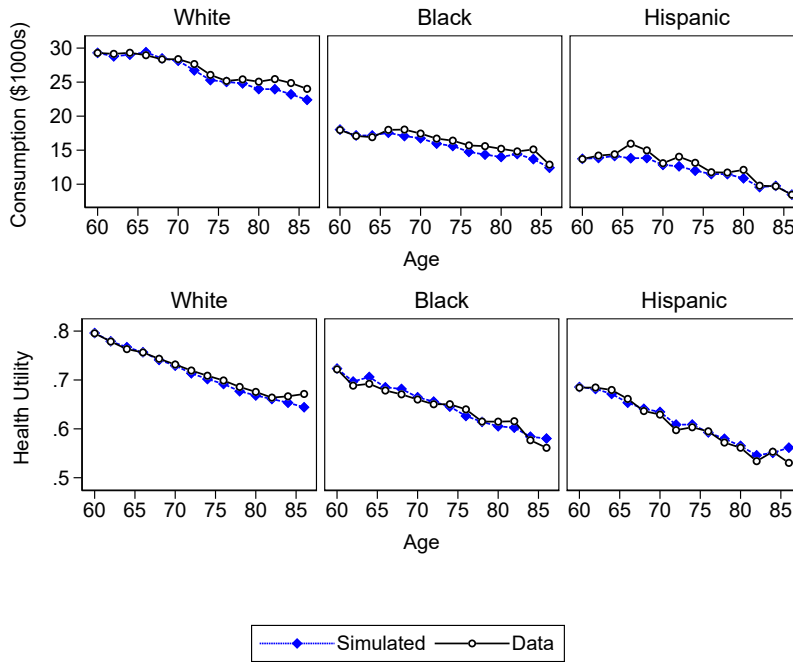
Table 4: Model estimates for morbidities

Variable	Hypertension		Diabetes		Cancer		Lung disease		Heart disease		Stroke		Psych		Arthritis	
	Coeff	SE	Coeff	SE	Coeff	SE	Coeff	SE	Coeff	SE	Coeff	SE	Coeff	SE	Coeff	SE
Lag Hyper			0.263	0.031	-0.033	0.037	0.079	0.038	0.153	0.031	0.112	0.038	0.124	0.036	0.079	0.031
Lag Diab			0.003	0.049	0.036	0.043	0.069	0.045	0.048	0.040	0.032	0.048	0.070	0.045	0.038	0.042
Lag Cancer			0.084	0.051	0.126	0.053	0.045	0.054	-0.060	0.046	-0.017	0.054	-0.064	0.056	0.043	0.047
Lag Lung			0.104	0.037	0.016	0.038	0.185	0.038	0.259	0.047	0.039	0.057	0.129	0.056	0.183	0.059
Lag Heart			-0.062	0.060	0.022	0.053	-0.047	0.058	0.090	0.049	0.154	0.037	0.071	0.039	0.085	0.039
Lag Stroke			0.073	0.047	-0.063	0.054	0.120	0.051	0.084	0.044	0.130	0.050	0.285	0.048	-0.077	0.059
Lag Psych			0.010	0.031	-0.038	0.033	0.146	0.034	0.058	0.028	0.002	0.035	0.104	0.033	0.277	0.050
Lag Arthritis			0.034	0.032	0.007	0.032	0.094	0.033	0.071	0.028	0.186	0.031	0.219	0.030		
Lag ADL			-0.017	0.030	-0.057	0.031	-0.083	0.030	-0.097	0.028	-0.116	0.030	-0.178	0.029	0.169	0.034
Lag Health 2			-0.023	0.032	-0.090	0.033	-0.162	0.033	-0.169	0.029	-0.216	0.030	-0.282	0.031	-0.059	0.035
Lag Health 3			-0.102	0.035	-0.121	0.035	-0.318	0.037	-0.258	0.031	-0.248	0.036	-0.382	0.036	-0.107	0.036
Lag Health 4			-0.224	0.043	-0.165	0.042	-0.435	0.051	-0.308	0.038	-0.331	0.047	-0.450	0.047	-0.153	0.038
Lag Health 5 (best)			0.020	0.031	0.073	0.036	-0.089	0.037	0.026	0.030	0.036	0.038	-0.064	0.035	-0.249	0.042
Lag2 Hyper			-0.083	0.049	-0.047	0.045	-0.136	0.048	0.103	0.041	0.093	0.050	-0.070	0.047	-0.029	0.044
Lag2 Diab			0.053	0.053	-0.010	0.057	0.030	0.057	0.076	0.049	-0.019	0.057	0.079	0.059	0.002	0.051
Lag2 Cancer			-0.066	0.056	-0.002	0.057	-0.087	0.040	-0.108	0.051	0.001	0.061	-0.016	0.061	-0.118	0.066
Lag2 Lung			-0.013	0.039	0.012	0.040	0.078	0.063	0.019	0.054	-0.003	0.039	-0.056	0.041	-0.027	0.042
Lag2 Heart			0.071	0.065	-0.036	0.058	0.010	0.053	-0.024	0.047	-0.027	0.053	-0.197	0.054	0.063	0.065
Lag2 Stroke			-0.082	0.050	0.054	0.057	0.010	0.053	0.031	0.028	-0.019	0.034	-0.008	0.032	-0.138	0.053
Lag2 Psych			-0.008	0.031	0.069	0.033	-0.049	0.033	0.031	0.028	-0.019	0.034	-0.065	0.032	-0.077	0.038
Lag2 Arthre			0.031	0.034	-0.023	0.034	-0.016	0.034	0.023	0.029	-0.097	0.032	-0.070	0.031	0.050	0.038
Lag2 ADL			-0.065	0.033	-0.044	0.033	-0.078	0.032	0.004	0.030	-0.063	0.032	-0.070	0.031	0.050	0.038
Lag2 Health 2			-0.009	0.035	-0.015	0.035	-0.131	0.034	-0.013	0.032	-0.060	0.034	-0.114	0.034	0.061	0.039
Lag2 Health 3			-0.016	0.036	-0.057	0.033	-0.179	0.037	-0.034	0.034	-0.052	0.037	-0.199	0.037	0.028	0.041
Lag2 Health 4			-0.028	0.037	-0.002	0.037	-0.179	0.037	-0.034	0.034	-0.098	0.046	-0.274	0.047	-0.041	0.044
Lag2 Health 5			-0.055	0.040	0.007	0.042	-0.284	0.049	-0.079	0.039	-0.098	0.046	-0.274	0.047	-0.041	0.044
Time			0.018	0.007	-0.001	0.007	0.004	0.008	-0.015	0.006	-0.032	0.007	-0.010	0.007	-0.031	0.007
2008+			-0.004	0.026	0.010	0.028	0.023	0.031	-0.004	0.025	0.042	0.030	-0.043	0.031	0.003	0.026
CODA			-0.018	0.037	-0.012	0.041	0.028	0.043	-0.021	0.034	0.017	0.037	0.069	0.040	-0.073	0.037
Early HRS			-0.077	0.050	-0.062	0.052	-0.013	0.059	0.001	0.046	-0.002	0.052	0.045	0.055	-0.067	0.051
Late HRS			-0.069	0.063	-0.085	0.066	0.021	0.075	0.037	0.058	0.016	0.066	0.070	0.070	0.032	0.063
War Babies			-0.065	0.078	-0.049	0.082	0.003	0.093	0.052	0.072	0.088	0.082	0.209	0.085	0.139	0.078
Boomers			-0.136	0.095	-0.110	0.100	-0.018	0.114	0.079	0.088	0.096	0.101	0.272	0.104	0.202	0.095
Mid Boomers			-0.232	0.111	0.020	0.121	0.025	0.135	0.140	0.104	0.147	0.120	0.261	0.123	0.208	0.111
Late Boomers			-0.255	0.131	-0.076	0.143	0.031	0.161	0.140	0.125	0.294	0.148	0.272	0.145	0.307	0.129
Black			0.209	0.020	0.106	0.019	-0.163	0.024	-0.160	0.019	0.031	0.022	-0.188	0.023	-0.036	0.019
Hispanic			0.104	0.023	0.254	0.023	-0.249	0.031	-0.232	0.024	-0.126	0.031	-0.030	0.028	-0.126	0.024
Female			0.006	0.015	-0.098	0.016	-0.033	0.018	-0.173	0.014	-0.064	0.018	0.138	0.017	0.159	0.014
HS grad			-0.025	0.018	-0.059	0.019	-0.104	0.021	0.007	0.017	0.022	0.021	-0.059	0.021	-0.052	0.019
Some college			-0.076	0.020	-0.064	0.021	-0.095	0.025	0.021	0.020	0.035	0.024	-0.002	0.024	-0.022	0.020
College grad			-0.106	0.023	-0.115	0.025	-0.215	0.030	-0.050	0.023	0.017	0.028	-0.016	0.028	-0.073	0.023
Lag Retired			-0.038	0.028	0.046	0.029	0.066	0.038	0.021	0.029	0.050	0.040	0.091	0.036	0.002	0.027
Lag2 Retired			0.028	0.028	-0.046	0.028	0.010	0.036	-0.005	0.028	0.028	0.037	-0.026	0.034	-0.015	0.027
Constant			-1.488	0.081	-1.923	0.086	-1.966	0.100	-1.666	0.078	-2.447	0.103	-1.781	0.090	-1.326	0.082

Notes: Multivariate probit results with dependent variable across columns. Regressions also include dummies for age, occupation, and census division.

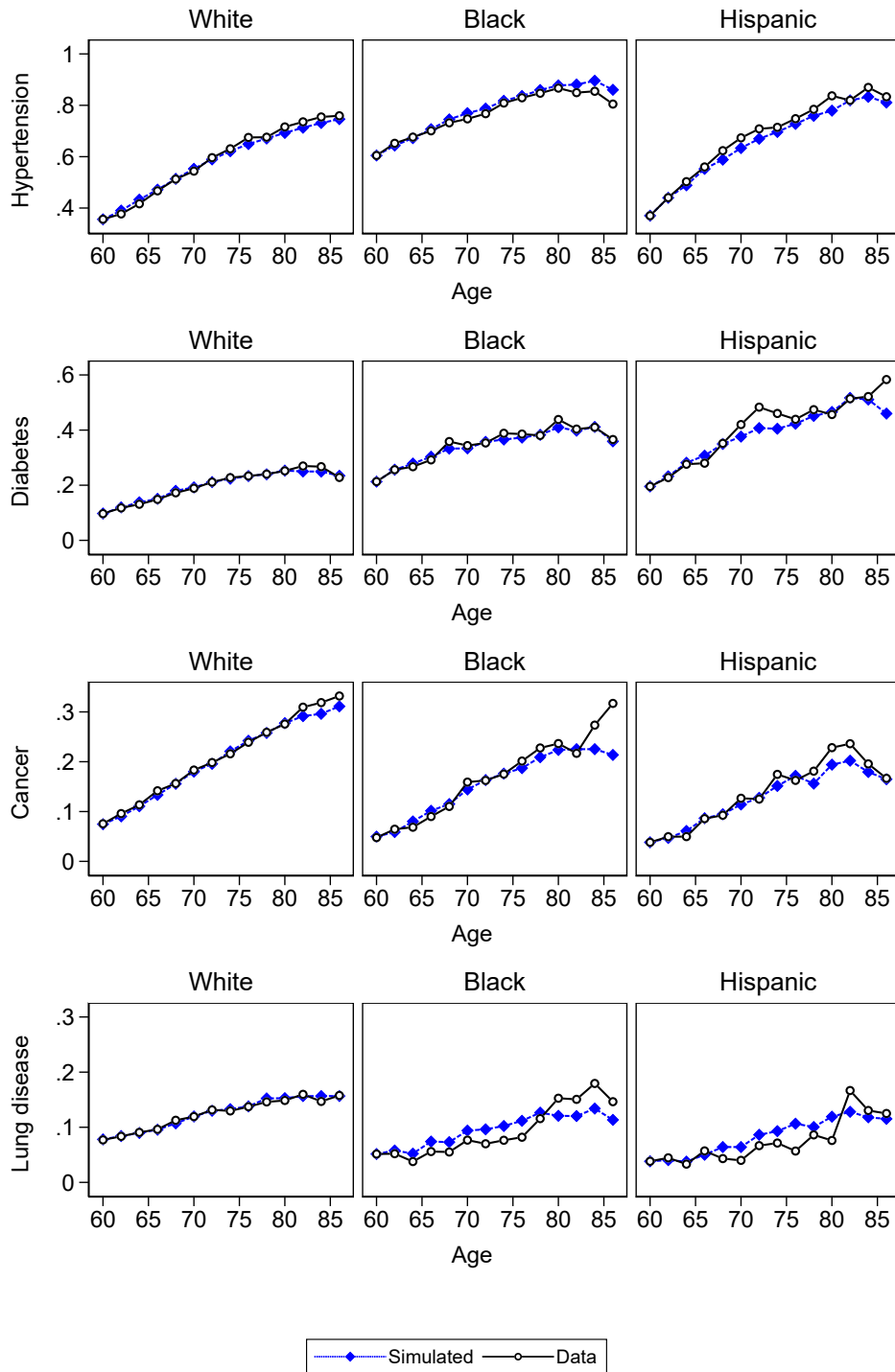
Table 5: Morbidity shock covariance matrix ( $\Sigma$ )

	Hyper	Diabetes	Cancer	Lung	Heart	Stroke	Psych	Arthritis	ADLs
Hyper	1.00	0.26	0.05	0.08	0.29	0.29	0.14	0.09	0.10
Diabetes	0.26	1.00	0.06	0.05	0.11	0.13	0.06	0.05	0.06
Cancer	0.05	0.06	1.00	0.12	0.05	0.06	0.10	0.06	0.14
Lung	0.08	0.05	0.12	1.00	0.22	0.11	0.17	0.09	0.18
Heart	0.29	0.11	0.05	0.22	1.00	0.28	0.16	0.10	0.14
Stroke	0.29	0.13	0.06	0.11	0.28	1.00	0.20	0.09	0.39
Psych	0.14	0.06	0.10	0.17	0.16	0.20	1.00	0.16	0.29
Arthritis	0.09	0.05	0.06	0.09	0.10	0.09	0.16	1.00	0.26
ADLs	0.10	0.06	0.14	0.18	0.14	0.39	0.29	0.26	1.00



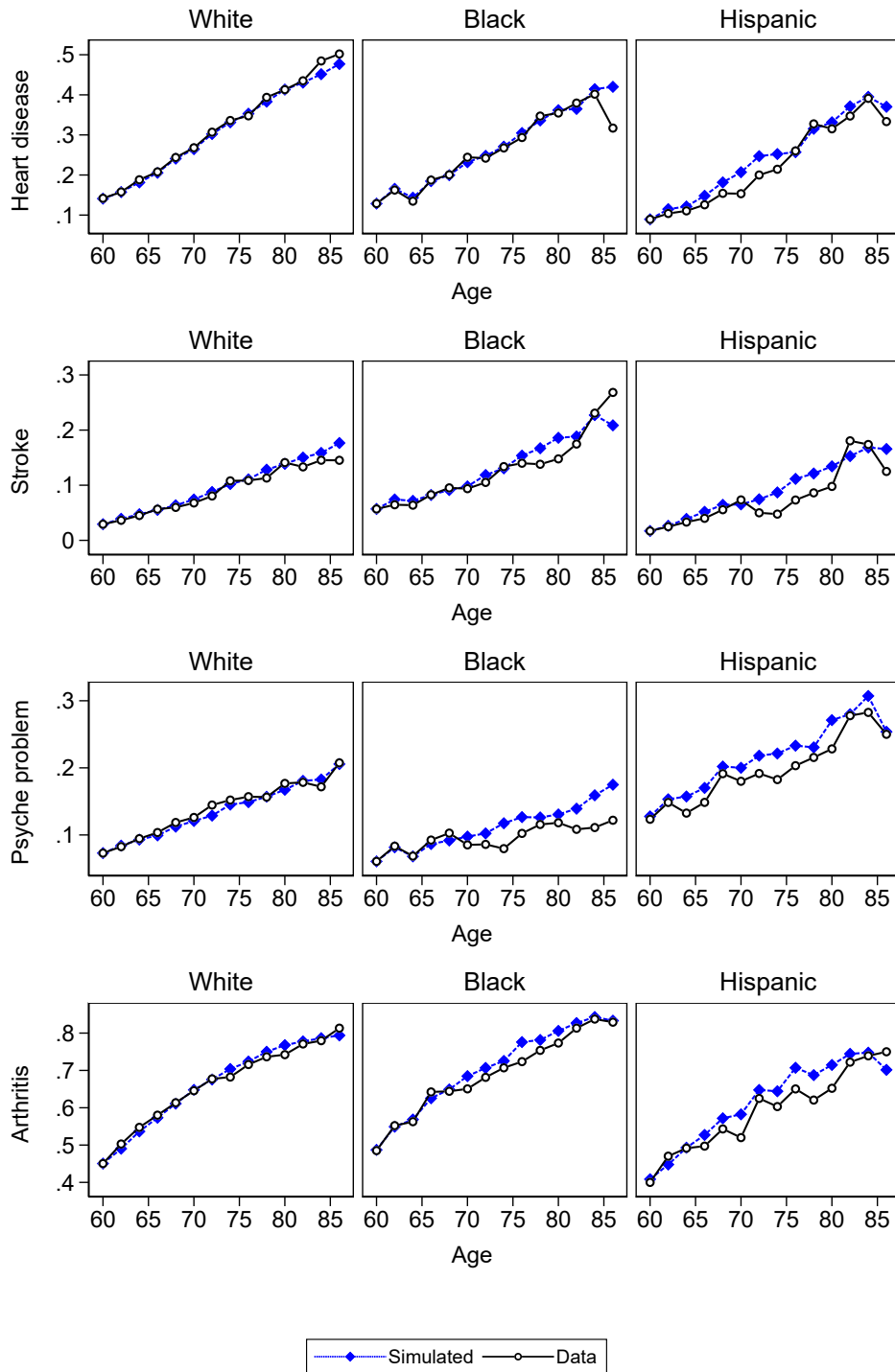
Notes: “Data” plots mean of all available data (inclusive of imputed missing values) in the EHRS cohort by two-year age interval. “Simulated” plots mean of expected simulated outcome for each observation in the data (i.e. the expected outcome for each person-year observation in the data).

Figure 1: Mean of life-cycle consumption and health utility profiles by race/ethnicity



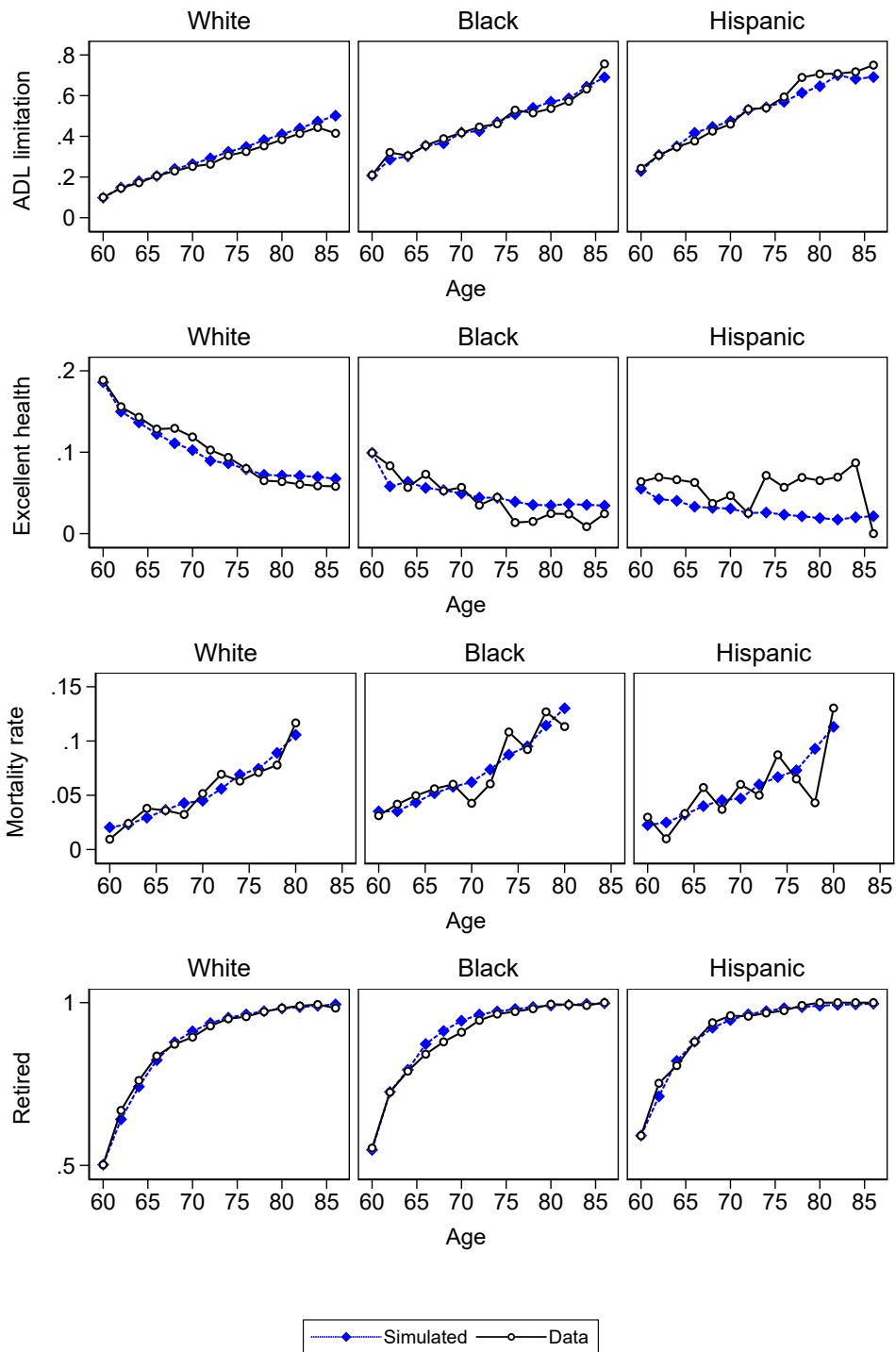
Notes: “Data” plots mean of all available data (inclusive of imputed missing values) in the EHRS cohort by two-year age interval. “Simulated” plots mean of expected simulated outcome for each observation in the data (i.e. the expected outcome for each person-year observation in the data).

Figure 2: Mean of life-cycle morbidity profiles by race/ethnicity



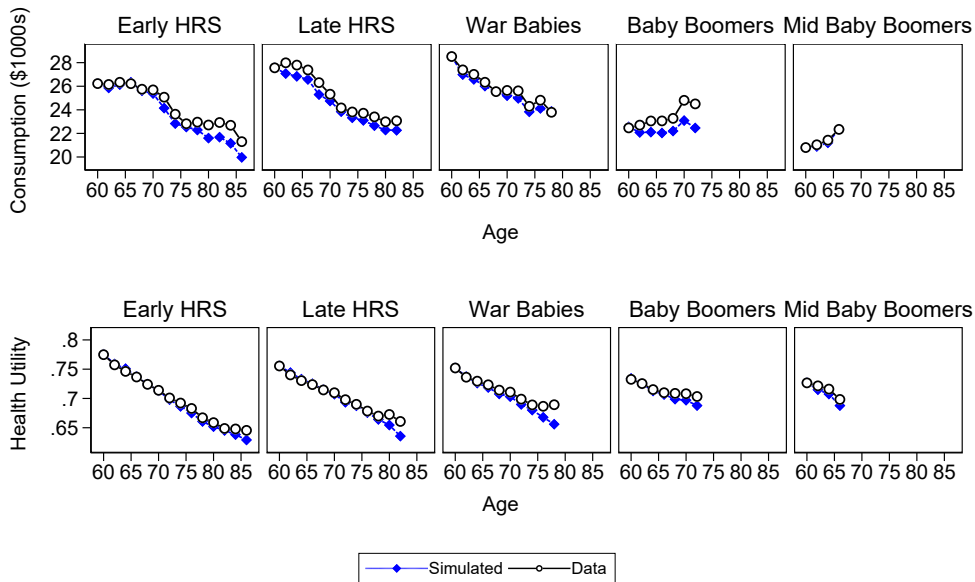
Notes: “Data” plots mean of all available data (inclusive of imputed missing values) in the EHR cohort by two-year age interval. “Simulated” plots mean of expected simulated outcome for each observation in the data (i.e. the expected outcome for each person-year observation in the data).

Figure 3: Mean of life-cycle morbidity profiles by race/ethnicity



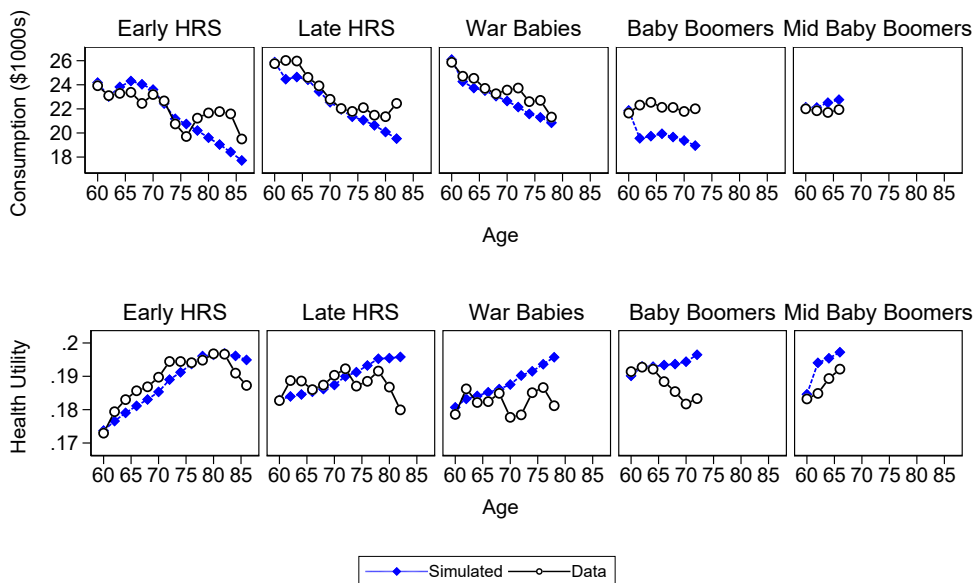
Notes: “Data” plots mean of all available data (inclusive of imputed missing values) in the EHRS cohort by two-year age interval. “Simulated” plots mean of expected simulated outcome for each observation in the data (i.e. the expected outcome for each person-year observation in the data).

Figure 4: Mean of life-cycle health, mortality, and retirement profiles by race/ethnicity



Notes: “Data” plots mean of all available data (inclusive of imputed missing values) in HRS by two-year age interval and cohort. “Simulated” plots mean of expected simulated outcome for each observation in the data (i.e. the expected outcome for each person-year observation in the data).

Figure 5: Mean of life-cycle consumption and health utility profiles by cohort



Notes: “Data” plots standard deviation of all available data (inclusive of imputed missing values) in HRS by two-year age interval and cohort. “Simulated” plots mean of standard deviations of simulated outcome (i.e. the mean of standard deviations calculated for each of the 5,000 simulation runs).

Figure 6: Standard deviation of consumption and health utility life-cycle profiles by cohort

## C Health utility weights

We obtain our health utility weights  $\omega$  from the Health Utilities Index Mark 3 (HUI3) instrument, which was administered to around 1,200 participants in the HRS in 2000. The HUI3 instrument was designed to produce cardinal utility scores on the standard utility scale of 0 (death) to 1 (best health) and has been widely used in studies on health utilities (Furlong et al., 1998; Feeny et al., 2002; Horsman et al., 2003). We use the HUI multi-attribute utility score ( $hui3ou$ ) for our analysis.

The HUI3 was conceptualized such that  $u(h_i) = HUI3_i \times u(h_{best})$  for individual  $i$  and general utility function  $u(\cdot)$ , where  $h_{best}$  refers to the best possible health state. For example, a year in the best health state is equivalent in utility to two years with  $HUI3 = 0.5$ . For our model, we adopt the approach of Miller and Bairoliya (2023) and assume that the HUI3 measures relative utility across health states *while holding consumption and leisure fixed*:

$$\omega h_i [\bar{u} + \log(c_i) + v(l_i)] = HUI3_i \times h_{best} [\bar{u} + \log(c_i) + v(l_i)].$$

This approach is consistent with the HUI3 instrument, as the interview script instructs participants to imagine themselves in the given health states while assuming that where they live, their income, and their friends and family remain constant. Given this assumption, the above equation simplifies to  $\omega h_i = HUI3_i$  when  $h_{best} = 1$ . We estimate the utility weights  $\omega$  by regressing the HUI3 utility score on self-rated health and all morbidity indicators. Estimated benchmark health utility weights are presented in Table 6.

Table 6: Estimated health utility weights ( $\gamma$ )

Measure	Weight	SE
Self-rated health		
Fair	0.229	0.026
Good	0.314	0.026
Very good	0.405	0.028
Excellent	0.421	0.031
Hypertension	0.004	0.012
Diabetes	-0.003	0.018
Cancer	0.009	0.017
Lung disease	-0.027	0.022
Heart disease	-0.031	0.015
Stroke	-0.077	0.022
Psych problem	-0.069	0.020
Arthritis	-0.062	0.013
Diff with ADL	-0.158	0.017
Constant	0.516	0.028

Notes: Results from regression of adjusted HUI3 score on self-rated health and morbidities. SE denotes standard error.  $R^2 = 0.17$ .  $N = 760$ .

While this approach is consistent with the interview instructions of the survey, other researchers have questioned whether respondents are fully capable of conceptualizing changes in health states without also considering changes in other aspects of life (Feeny et al., 2018). For instance, respondents may have considered changes in consumption and leisure along with changes in health. In such cases, the appropriate representation of the HUI3 instrument would be as follows:

$$\gamma h [\bar{u} + \log(c) + v(l)] = HUI3 \times h_{best} [\bar{u} + \log(c_{best}) + v(l_{best})].$$

Rearranging terms and setting  $h_{best} = 1$  yields:

$$\gamma h = HUI3 \frac{\bar{u} + \log(c_{best}) + v(l_{best})}{\bar{u} + \log(c) + v(l)}. \quad (1)$$

However, Miller and Bairoliya (2023) note that this formulation poses a problem because we do not observe the counterfactual consumption and leisure bundles that would be realized in the best health state. Nevertheless, we have already developed an independent forecasting model that enables us to predict the expected value for  $c_{best}$  and  $l_{best}$  for each individual in the sample. Armed with these predictions, we calculated the right-hand side of (1) for each HUI3 respondent. We then regressed this value on self-rated health and all morbidity indicators to obtain alternate utility weights  $\gamma$  (see results in Table 7). We used these alternative utility weights in our robustness exercises.

Table 7: Estimated alternate health utility weights ( $\gamma$ )

Measure	Weight	SE
Self-rated health		
Fair	0.267	0.035
Good	0.336	0.035
Very good	0.417	0.037
Excellent	0.408	0.042
Hypertension	-0.002	0.017
Diabetes	0.015	0.024
Cancer	0.002	0.023
Lung disease	-0.037	0.029
Heart disease	-0.045	0.021
Stroke	-0.054	0.030
Psych problem	-0.059	0.028
Arthritis	-0.061	0.017
Diff with ADL	-0.139	0.022
Constant	0.507	0.038

Notes: Results from regression of adjusted HUI3 score on self-rated health and morbidities. SE denotes standard error.  $R^2 = 0.17$ . N = 760.



## D Additional welfare results

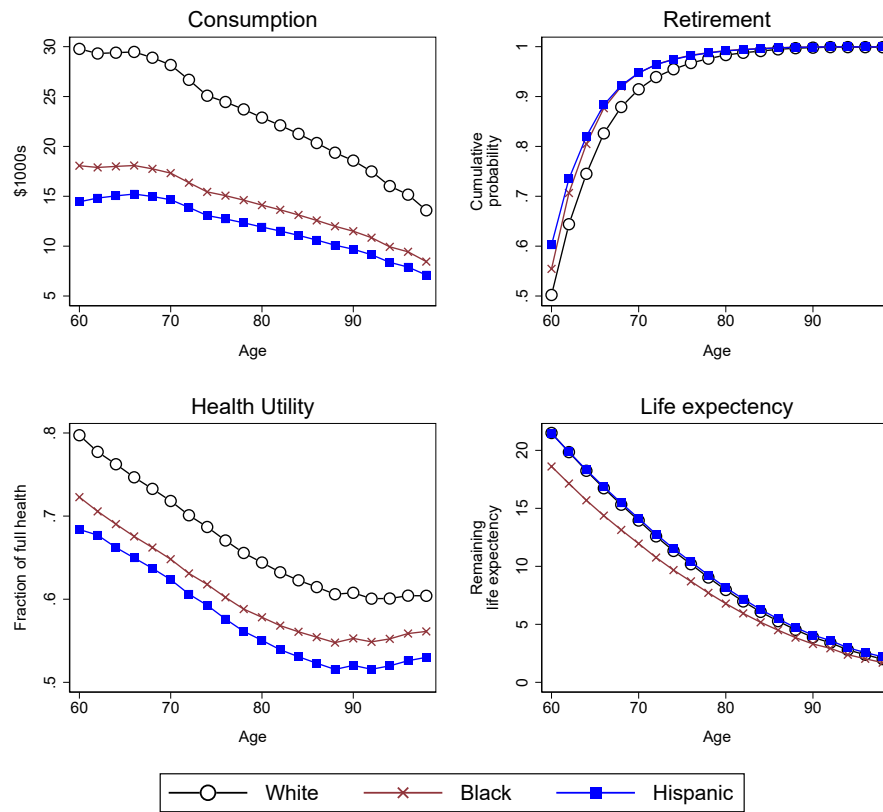


Figure 7: Average life cycle profiles by race/ethnicity

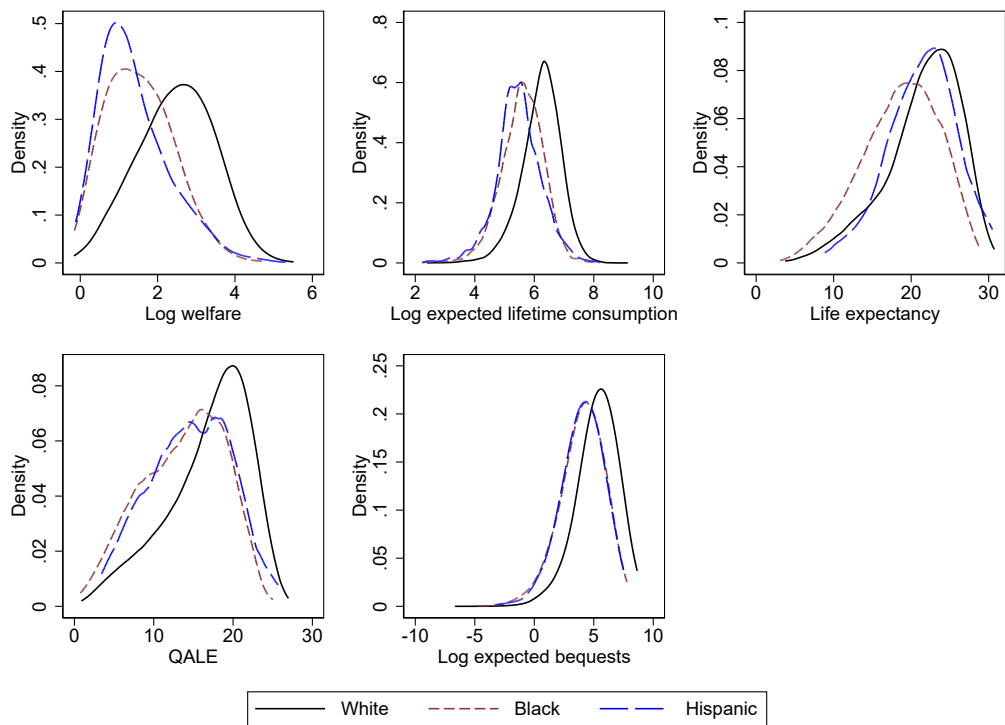
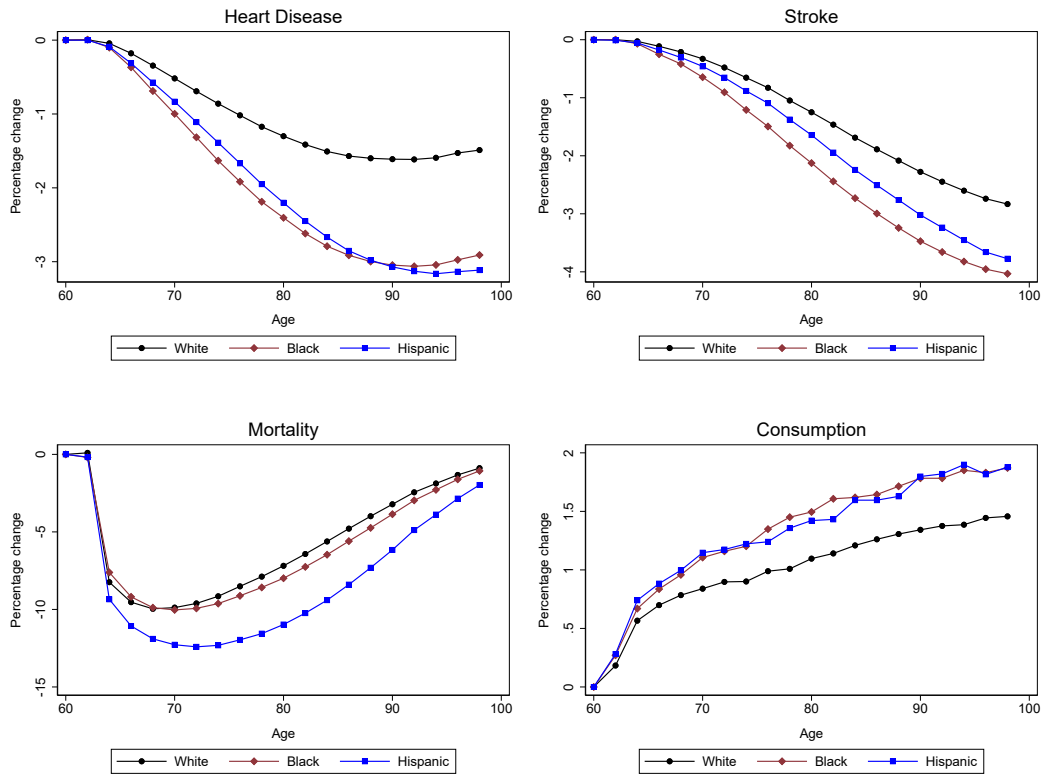


Figure 8: Distribution of welfare, consumption, life expectancy, and bequests by race/ethnicity



*Notes:* Results plot percentage difference in expected outcomes with the exogenous elimination of diabetes after age sixty relative to baseline. Sample includes all individuals in the simulation sample from the EHRS cohort. Expected outcomes are conditional on survival.

Figure 9: Impulse response to elimination of diabetes after age 60

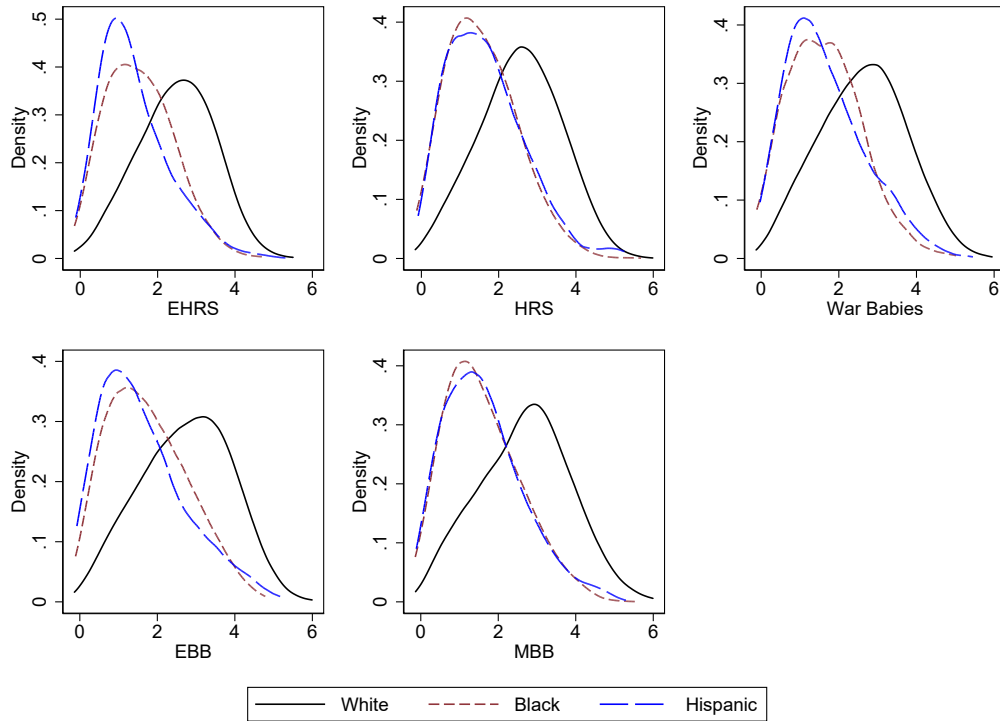


Figure 10: Distribution of log welfare by race and cohort

Table 8: Cumulatively adjusted welfare ratios by cohort

Measure	Black-White Ratio					Hispanic-White Ratio				
	EHRS	LHRS	WB	EBB	MBB	EHRS	LHRS	WB	EBB	MBB
Consumption	0.613	0.593	0.602	0.599	0.540	0.516	0.560	0.525	0.513	0.473
Leisure	0.623	0.604	0.620	0.619	0.558	0.526	0.566	0.537	0.523	0.486
Life Exp.	0.501	0.479	0.466	0.462	0.402	0.524	0.610	0.516	0.479	0.435
Health	0.468	0.436	0.417	0.430	0.366	0.446	0.546	0.454	0.413	0.375
Bequests ( $\lambda$ )	0.378	0.356	0.347	0.358	0.293	0.369	0.467	0.404	0.374	0.328

Notes: Estimates using base year respondent analysis weights.

## References

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